# Simple least squares

Summary

- → Model form:  $y = a_0 + a_1 x + e$
- →  $S_r = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i a_0 a_1 x_i)^2$  becomes minimizes where  $\frac{\partial S_r}{\partial a_0} = 0 \& \frac{\partial S_r}{\partial a_1} = 0$ .

→ Rearranging and solving for  $a_0$  and  $a_1$ 

$$na_{0} + \left(\sum x_{i}\right)a_{1} = \sum y_{i} \qquad \left(\sum x_{i}\right)a_{0} + \left(\sum x_{i}^{2}\right)a_{1} = \sum x_{i}y_{i}$$
$$a_{1} = \frac{n\sum x_{i}y_{i} - \sum x_{i}\sum y_{i}}{n\sum x_{i}^{2} - \left(\sum x_{i}\right)^{2}} \qquad a_{0} = \overline{y} - a_{1}\overline{x}$$

Question: what if our model we want to find is non-linear?

Ex. Activation energy in rate constant

$$k = k_0 e^{-E/RT}$$

→ Linearize !

# Linearization

- Want to model non-linear relationships between independent (x) and dependent (y) variables.
  - 1. Make a simple linear model through a suitable transformation.

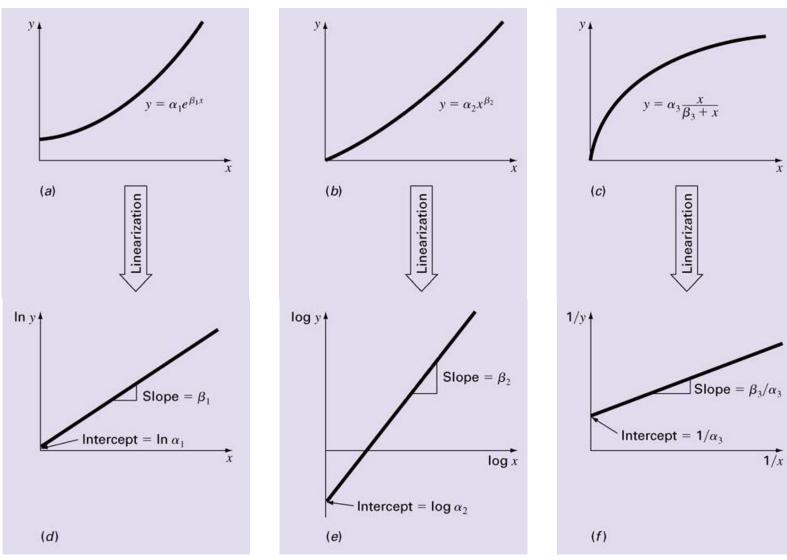
$$y = f(x) + e \rightarrow y = a_0 + a_1 x + e$$

2. Use previous results (simple least squares)

$$a_{1} = \frac{n \sum x_{i} y_{i} - \sum x_{i} \sum y_{i}}{n \sum x_{i}^{2} - \left(\sum x_{i}\right)^{2}} \qquad a_{0} = \overline{y} - a_{1} \overline{x}$$

\*Caution: transformation also changes P.D.F of variables (and errors) We will discuss about this in model assessment.

Linearization (Cont.)



#### 공정 모형 및 해석, 유준 02010

## Polynomial regression

✤ For quadratic form

$$y = a_0 + a_1 x + a_2 x^2 + e$$

✤ Sum of squares

$$S_r = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n \left( y_i - a_0 - a_1 x_i - a_2 x_i^2 \right)^2$$

Again,  $S_r$  has a parabolic shape w.r.t  $a_0$ ,  $a_1$ , and  $a_2$ . with plus signs of  $a_0^2$ ,  $a_1^2$ , and  $a_2^2$ .

$$\frac{\partial S_r}{\partial a_0} = -2\sum (y_i - a_0 - a_1 x_i - a_2 x_i^2) = 0$$
$$\frac{\partial S_r}{\partial a_1} = -2\sum x_i (y_i - a_0 - a_1 x_i - a_2 x_i^2) = 0$$
$$\frac{\partial S_r}{\partial a_2} = -2\sum x_i^2 (y_i - a_0 - a_1 x_i - a_2 x_i^2) = 0$$

공정 모형 및 해석, 유준 02010

2010-11-03

# Polynomial regression (Cont.)

✤ Rearranging the previous equations gives

$$(n)a_{0} + (\sum x_{i})a_{1} + (\sum x_{i}^{2})a_{2} = \sum y_{i} (\sum x_{i})a_{0} + (\sum x_{i}^{2})a_{1} + (\sum x_{i}^{3})a_{2} = \sum x_{i}y_{i} (\sum x_{i}^{2})a_{0} + (\sum x_{i}^{3})a_{1} + (\sum x_{i}^{4})a_{2} = \sum x_{i}^{2}y_{i}$$
 
$$(n \sum x_{i} \sum x_{i}^{2} \sum x_{i}^{2} \sum x_{i}^{2}) (a_{0}) (\sum x_{i}^{2} \sum x_{i}^{2} \sum x_{i}^{3}) (a_{1}) (a_{1}) = (\sum y_{i}) (a_{1}) (\sum x_{i}^{2}) (a_{0}) (a_{1}) = (\sum x_{i}y_{i}) (a_{1}) (a_{1}) (a_{1}) (a_{1}) (a_{1}) = (\sum x_{i}y_{i}) (a_{1}) (a_{1})$$

the above equations can be solved easily. (three unknowns and three equations.)

For general polynomials

$$y = a_0 + a_1 x + a_2 x^2 + \dots + a_m x^m + e$$

→ From the results of two cases  $(y = a_0 + a_1x \& y = a_0 + a_1x + a_2x^2)$ 

$$\frac{\partial S_r}{\partial a_0} = \frac{\partial S_r}{\partial a_1} = \dots = \frac{\partial S_r}{\partial a_m} = 0$$

we need to solve (m+1) linear algebraic equations for (m+1) parameters.

### Multiple least squares

Consider when there are more than two independent variables, x₁, x₂,
 ..., x<sub>m</sub>. → regression plane.

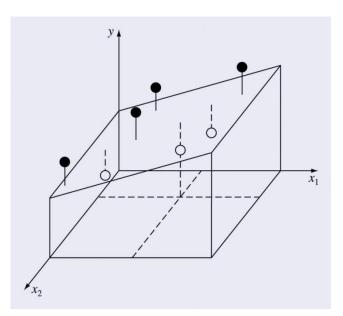
$$y = a_0 + a_1 x_1 + a_2 x_2 + \dots + a_m x_m + e$$

→ For 2-D case,  $y = a_0 + a_1 x_1 + a_2 x_2$ .

→ Again,  $S_r$  has a parabolic shape w.r.t  $a_0$ ,  $a_1$ .

$$S_r = \sum (y_i - a_0 - a_1 x_{1,i} - a_2 x_{2,i})^2$$

$$\frac{\partial S_r}{\partial a_0} = -2\sum (y_i - a_0 - a_1 x_{1,i} - a_2 x_{2,i}) = 0$$
$$\frac{\partial S_r}{\partial a_1} = -2\sum x_{1,i} (y_i - a_0 - a_1 x_{1,i} - a_2 x_{2,i}) = 0$$
$$\frac{\partial S_r}{\partial a_2} = -2\sum x_{2,i} (y_i - a_0 - a_1 x_{1,i} - a_2 x_{2,i}) = 0$$



# Multiple least squares (Cont.)

→ Rearranging and solve for  $a_0$ ,  $a_1$  and  $a_2$  gives

$$\begin{pmatrix} n & \sum x_{1,i} & \sum x_{2,i} \\ \sum x_{1,i} & \sum x_{1,i}^2 & \sum x_{1,i}x_{2,i} \\ \sum x_{2,i} & \sum x_{1,i}x_{2,i} & \sum x_{2,i}^2 \end{pmatrix} \begin{cases} a_1 \\ a_2 \\ a_3 \end{cases} = \begin{cases} \sum y_i \\ \sum x_{1,i}y_i \\ \sum x_{2,i}y_i \end{cases}$$

For an m-dimensional plane,

$$y = a_0 + a_1 x_1 + a_2 x_2 + \dots + a_m x_m + e$$

✤ Same as in general polynomials,

$$\frac{\partial S_r}{\partial a_0} = \frac{\partial S_r}{\partial a_1} = \dots = \frac{\partial S_r}{\partial a_m} = 0$$

we need to solve (m+1) linear algebraic equations for (m+1) parameters.

# General least squares

 The following form includes all cases (simple least squares, polynomial regression, multiple regression)

$$y = a_0 z_0 + a_1 z_1 + a_2 z_2 + \dots + a_m z_m + e$$
  
where  $z_0, z_1, \dots, z_m$  :  $m + 1$  different functions

Ex. Simple and multiple least squares

$$Z_0 = 1, Z_1 = x_1, Z_2 = x_2, \cdots, Z_m = x_m$$

polynomial regression

$$Z_0 = x^0 = 1, Z_1 = x^1, Z_2 = x^2, \dots, Z_m = x^m$$

✤ Same as before,

$$\frac{\partial S_r}{\partial a_0} = \frac{\partial S_r}{\partial a_1} = \dots = \frac{\partial S_r}{\partial a_m} = 0$$

we need to solve (m+1) linear algebraic equations for (m+1) parameters.

2010-11-03

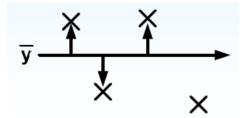
# Quantification of errors

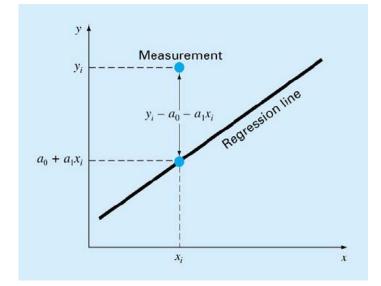
$$S_t = \sum (y_i - \overline{y})^2$$

$$S_r = \sum e_i^2$$
  
=  $\sum (y_i - a_0 z_{0,i} - a_1 z_{1,i} - \dots - a_m z_{m,i})^2$ 

### Total sum of squares around the mean for the response variable, *y*

Sum of squares of residuals around the regression line

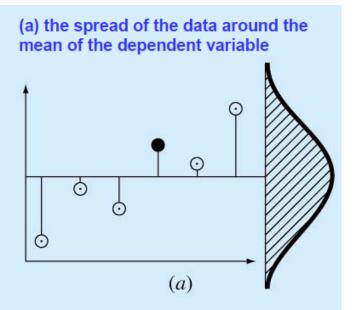




## Quantification of errors (Cont.)

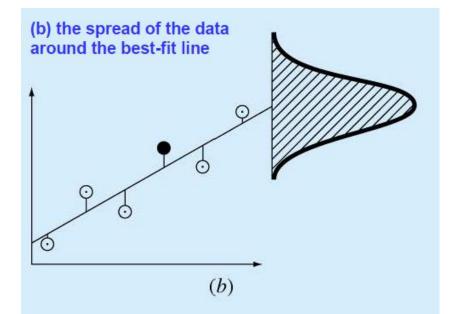
$$S_{y} = \sqrt{\frac{1}{n-1}\sum(y_{i} - \bar{y})^{2}} = \sqrt{\frac{S_{t}}{n-1}}$$

Standard deviation of y



$$S_{y/x} = \sqrt{\frac{S_r}{n - (m+1)}}$$

Standard error of predicted y  $\rightarrow$  quantify appropriateness of regression



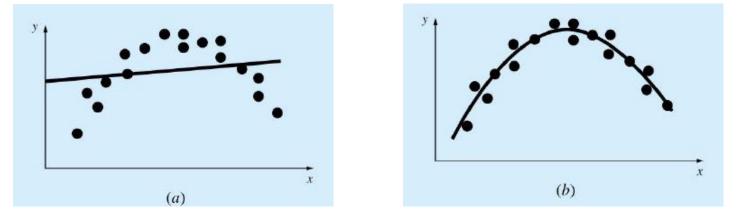
# Quantification of errors (Cont.)

Coefficients of determination, R<sup>2</sup>

$$R^2 = \sqrt{\frac{S_t - S_r}{S_t}}$$

The amount of variability in the data explained by the regression model.

 $R^2 = 1$  when  $S_r = 0$ : perfect fit (a regression curve passes through data points)  $R^2 = 0$  when  $S_r = S_t$ : as bad as doing nothing

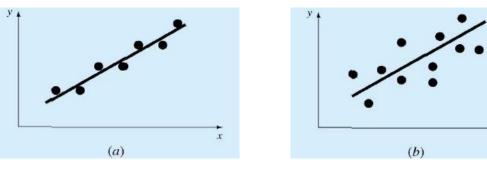


It is evident from the figures that a parabola is adequate.  $R^2$  of (b) is higher than that of (a)

# Quantification of errors (Cont.)

- → Warning! : R<sup>2</sup> ≈ 1 does not guarantee that the model is adequate, nor the model will predict new data well.
  - It is possible to force R<sup>2</sup> to be one by adding as many terms as there are observations.
  - $S_r$  can be big when variance of random error is large.

(Usual assumption on error is that error is random is unpredictable)



Practice using Minitab

- (1) Wind tunnel example with higher polynomials
- (2) Simple regression with increasing random noise

# Confidence intervals - coefficients

• Coefficients in the regression model have confidence interval.

 $y = a_0 z_0 + a_1 z_1 + a_2 z_2 + \dots + a_m z_m + e$ 

♦ Why? They are also statistics like x̄ & s. That is, they are numerical quantities calculated in a sample (not entire population). They are estimated values of parameters.

$$statistic \pm A \times \sigma_{statistic}$$

Value that depends on P.D.F of the statistic & confidence level  $\boldsymbol{\alpha}$ 

Standard error of the statistic

statistic	Α	$\sigma_{ m statistic}$
$\overline{x}$	$z_{\alpha/2}$	$\sigma_{_x}/\sqrt{n}$
$\overline{x}$	$t_{v,\alpha/2}$	$s_x/\sqrt{n}$

\* The standard error of a statistic is the standard deviation of the sampling distribution of that statistic

Statistic that we want to find

its confidence interval

#### 공정 모형 및 해석, 유준 02010

# Confidence intervals – coefficients (cont.)

Matrix representation of GLS

$$y = a_0 z_0 + a_1 z_1 + a_2 z_2 + \dots + a_m z_m + e$$

$$\mathbf{y} = \mathbf{Z}\mathbf{a} + \mathbf{e}$$

matrix of the calculated values of the basis functions at the measured values of the independent variable
observed valued of the dependent variable
unknown coefficients
residuals

$$\mathbf{Z} = \begin{bmatrix} Z_{01} & Z_{11} & \cdots & Z_{m1} \\ Z_{02} & Z_{12} & \cdots & Z_{m2} \\ \vdots & \vdots & \ddots & \vdots \\ Z_{0n} & Z_{1n} & \cdots & Z_{mn} \end{bmatrix} \quad \begin{aligned} \mathbf{y}^T &= \begin{bmatrix} y_1 & y_2 & \cdots & y_n \end{bmatrix} \\ \mathbf{a}^T &= \begin{bmatrix} a_0 & a_1 & \cdots & a_m \end{bmatrix} \\ \mathbf{e}^T &= \begin{bmatrix} e_1 & e_2 & \cdots & e_n \end{bmatrix} \end{aligned}$$

m+1: number of coefficients
n: number of data points

# Confidence intervals – coefficients (Cont.)

♦ Example

Fitting quadratic polynomials to five data points

$$y = a_0 + a_1 x + a_2 x^2 + e$$

$$\mathbf{y} = \mathbf{Z}\mathbf{a} + \mathbf{e}$$

$$\begin{bmatrix} 1.0 \\ 0.5 \\ 0.0 \\ 0.5 \\ 0.5 \\ 2.0 \end{bmatrix} \begin{bmatrix} 1 & -1.0 & 1.0 \\ 1 & -0.5 & 0.25 \\ 1 & 0.0 & 0.0 \\ 1 & 0.5 & 0.25 \\ 1 & 1.0 & 1.0 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ a_2 \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \\ e_5 \end{bmatrix}$$

Three unknowns Five equations

#### Can you solve this problem?

# Confidence intervals – coefficients (Cont.)

✤ Solutions

y = Za + e

Sum of squares of errors

1. LU decomposition or other methods to solve L.A.E

 $(\mathbf{Z}^T\mathbf{Z})\mathbf{a} = \mathbf{Z}^T\mathbf{y} \implies \mathbf{A}\mathbf{x} = \mathbf{b}''$ 

2. Matrix inversion

$$(\mathbf{Z}^T \mathbf{Z})\mathbf{a} = \mathbf{Z}^T \mathbf{y}$$
  $\Rightarrow \mathbf{a} = (\mathbf{Z}^T \mathbf{Z})^{-1} \mathbf{Z}^T \mathbf{y}$ 

computationally not efficient, but statistically useful

# Confidence intervals – coefficients (Cont.)

Matrix inversion approach

 $\mathbf{a} = \left(\mathbf{Z}^T \mathbf{Z}\right)^{-1} \mathbf{Z}^T \mathbf{y}$ 

Denote  $Z_{ii}^{-1}$  as the diagonal element of  $(\mathbf{Z}^T \mathbf{Z})^{-1}$ Confidence interval of estimated coefficients

$$a_{i-1} \pm t_{n-(m+1),\alpha/2} \sqrt{S_{y/x}^2 Z_{ii}^{-1}}$$

$$t_{n-(m+1),\alpha/2}$$
 Student t statistics  
 $S_{y/x} = \sqrt{\frac{S_r}{n-(m+1)}}$  Standard error of e

rd error of estimate

### What if confidence intervals contain zero?